

Weierstrass Institute for Applied Analysis and Stochastics



# Connectivity networks in neuroscience - construction and analysis

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## The human brain







(Wikimedia)

Figure: John A Beal, PhD Dep't. of Cellular Biology & Anatomy, Louisiana State University Health Sciences Center Shreveport



CSF (Cere-brospinal fluid)







Figure: Blausen.com staff (2014). "Medical gallery of Blausen Medical 2014". WikiJournal of Medicine 1 (2). DOI:10.15347/wjm/2014.010. ISSN 2002-4436.

Lobes, Gyri (ridge on the cerebral cortex), Sulci (depression or groove in the cerebral cortex)

White Matter: Fiber bundles connecting cortical areas



#### **Functional areas**







## **Brodman Areas:**

- $\sim 50$  cortex regions defined based on cytoarchitectural organization of neurons
- regions have been correlated to cortical functions

Brodmann K (1909). "Vergleichende Lokalisationslehre der Grosshirnrinde". Leipzig: Johann Ambrosius Barth

Figure: Mark Dow. Research Assistant Brain Development Lab, University of Oregon.

http://lcni.uoregon.edu/ dow/Space\_software/renderings.html



## **Brain template**



Talaraich atlas: single subject Talairach et al. Co-planar stereotaxic atlas of the human brain. Thieme, New York. (1988)

Problem: Subject variability



## **Brain template**



Talaraich atlas: single subject Talairach et al. Co-planar stereotaxic atlas of the human brain. Thieme, New York. (1988)

Problem: Subject variability

MNI (Montreal Neurological Institute) template (ICBM152): Average of 152 MRI scans matched by affine transform (9 parameters) Maintained by the International Consortium for Brain Mapping



Orthographic view of MNI template





Havard-Oxford atlas (FSL) : 48 cortical + 21 subcortical regions, 37 subjects







Probabilistic atlas used in FSL (Desikan et al., (2006). NeuroImage, 31(3):968-80.)





Partially addresses subject variability







- Term covers a number of minimally invasive techniques to study the brain
- used to characterize structure, function and diagnostic of diseases
- contribute to understanding interactions between mind (decisions, emotions), brain and body

Two categories

Structural neuroimaging

Functional neuroimaging





- Term covers a number of minimally invasive techniques to study the brain
- used to characterize structure, function and diagnostic of diseases
- contribute to understanding interactions between mind (decisions, emotions), brain and body

Two categories with modalities

# Structural neuroimaging

- Computed tomography (CT)
- Positron emission tomography (PET)
- Magnetic resonance imaging (MRI)
- Diffusion weighted magnetic resonance imaging (dMRI/DWI)

# Functional neuroimaging

- Electroencephalography (EEG)
- Magnetoencephalography (MEG)
- Positron emission tomography (PET)
- functional magnetic resonance imaging (fMRI)





Describes the interaction of cortical brain regions

- Functional connectivity: characterises the simultaneous function of different brain regions
- Structural (anatomic) connectivity: describe the anatomical connection of functional brain regions (nodes) by white matter fiber tracks
- Effective connectivity: describe the causal interaction of functional brain regions by directed graphs

All require definition of nodes (functional regions) by some methods:

- use of anatomic information ( cortical thickness, myelination )
- functional regions identified by fMRI experiments
- default networks identified by resting state fMRI experiments









#### Magnetic resonance imaging (MRI)





Figure: Kasuga Huang (Wikimedia)



Figure: Franz Wilhelmstötter (Wikimedia)



Fig. 2.7 Free Induction Decay (F1D) following a single 90° r.f. pulse. The real and imaginary parts of the signal correspond to the in-phase and quadrature receiver outputs. The signal is depicted with receiver phase  $\phi = 0$  and, on complex Fourier transformation, gives real absorption and imaginary dispersion spectra at the offset frequency,  $\Delta \omega = \omega_0 - \omega$ .

From O. Friman "Adaptive Analysis of Functional MRI Data", PhD Thesis, 2003



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Figure: Kasuga Huang (Wikimedia)



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## **MR contrasts**





T1-weighted image (orthographic view), longitudinal relaxation



T2-weighted image (orthographic view), transverse relaxation

<sup>0</sup>Data provided by the Human Connectome Project, WU-Minn Consortium (Principal Investigators: David Van Essen and Kamil Ugurbil; 1U54MH091657) funded by the 16 NIH Institutes and Centers that support the NIH Blueprint for Neuroscience Research; and by the McDonnell Center for Systems Neuroscience at Washington University.

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## From K-space to image





- 24-32 receiver coils
- Acquisition protocol and reconstruction method cause (non-local) spatial correlation
- Signal distribution depends on the reconstruction method (SENSE, GRAPPA, SoS, ...)





# Structural MR images:

- high spatial resolution
- offer contrast between tissue types (cortex <-> white matter)
- no temporal information







#### Structural MR images:

- high spatial resolution
- offer contrast between tissue types (cortex <-> white matter)
- no temporal information



#### Functional MR images

- Iower spatial resolution
- Iower image contrast
- temporal resolution
- signal changing with experimental tasks







Lnibniz

Uses the Blood Oxygenation Level Dependent (BOLD) contrast

- Active neurons need oxygen!
- Change of magnetic properties due to oxygenation.
- Measure the ratio of oxygenated to deoxygenated hemoglobin
- Local signal changes over time due to brain function

Experiment:

Measurement of fast time series of the brain under stimulus

## Indirect measurement:

- Measures oxygen consumption of active neurons
- Signal changes are delayed in time
- Convolution with hemodynamic response function
- Limited spatial resolution by vascular architecture





# fMRI data = 3D + T



# About fMRI data

- Time series of 3D data
- Spatial resolution:
   1-4 mm
- Temporal resolution: 1-3 sec
- Search for locations, were a BOLD signal can be found!
- Problem: noise
- Problem: multiple test problem





# Realignment/Registration:

- Corrections for head movement
- Rigid or affine transformation

# Slice time correction:

- Adjust for slice recording at different times

# Normalization:

- Mapping to a standard space (Talaraich, MNI)
- Comparability between subjects in group studies
- Cortex segmentation based on corresponding anatomic images
- Spatial smoothing





#### Hemodynamic response function:



Parametric model:

$$h(t) = \left(\frac{t}{d_1}\right)^{a_1} \exp\left(-\frac{t-d_1}{b_1}\right)$$
$$-c\left(\frac{t}{d_2}\right)^{a_2} \exp\left(-\frac{t-d_2}{b_2}\right)$$

Time delay modeled by including the derivative of h

Figure: M. Lindquist, J. Hopkins Univ., Talk at SAMSI 2015 Study: Lindquist et al., Journal of Magnetic Resonance 2008

Spatially varying form and latency

Expected BOLD response: Convolution between stimulus and hemodynamic response





Linear model:

$$Y_i = X\beta_i + \varepsilon_i$$

Data 
$$Y_i = (Y_{it})$$

■ Design  $X_i = (x_{itk})$ , *i* - voxel, *t*-time, k = 1, K - components

Error  $\varepsilon_i = (\varepsilon_{it})$ ,  $\mathbf{E} \varepsilon_{it} = 0$ ,  $\mathbf{E} \varepsilon_{it}^2 = \sigma_t^2$ ,  $\mathbf{Cov}(\varepsilon_{it}, \varepsilon_{i(t-j)}) = \delta_{ij}$ , usually AR(1) or AR(2)

Components include:

- Expected bold responses to stimuli
- drift components for magnetic field inhomogeneity (polynomial)
- confounding (physiological) effects (respiration, cardiac cycle, ...)
- parameters from motion correction

Prewhitening: Transform model such that errors are approx. uncorrelated

$$\tilde{Y}_i = \tilde{X}_i \beta_i + \tilde{\varepsilon}_i, \quad \tilde{Y}_i = A_i Y_i, \quad \tilde{X}_i = A_i X, \quad \tilde{\varepsilon}_i = A_i \varepsilon_i,$$





## Learning paradigm:







## Stimulus components in design matrix:











Interest in contrast

$$\gamma = c^T \beta$$

and testing

$$H: \gamma_i = 0$$
 against  $A: \gamma_i <> 0$ 

to determine active brain regions associated with the contrast

Estimate AR(k) parameters from residuals in linear model  $Y_i = X\beta_i + \varepsilon_i$ 

- Spatially smooth AR(k) parameters
- Prewhitening using  $\hat{A}_i$  obtained from smoothed AR(k) parameters

Estimation of  $\beta_i$ :

$$\hat{\beta}_i = \left(\tilde{X}_i^T \tilde{X}_i\right)^{-1} \tilde{X}_i \tilde{Y}_i$$

Estimate covariance  $\hat{\Sigma}_i$  of  $\hat{\beta}_i$  from prewhitened model

Define test statistics (t-distributed)

$$S_i = \frac{c^T \hat{\beta}_i}{(c^T \Sigma_i c)^{1/2}}$$





- Simultaneous tests in N = 10000 (Cortex) 100000 (Brain) voxel
- using *t*-thresholds at significance level  $\alpha$  gives  $\approx \alpha N$  false positives.
- Adjustment for multiple testing by Bonferroni leads to high thresholds
- Multiplicity adjustment leads to low sensitivity
- Alternative: False Discovery Rate (Benjamini & Hochberg 1995)

Control of proportion of false positives within detected signals

ignores spatial extend of regions of interest



voxelwise decision using thresholds adjusted for multiple testing





Regions of activation have a spatial extend

Smoothing the observed images with a (Gaussian) kernel with bandwidth h

$$\bar{Y}_{it} = \sum_{j} K(\frac{||i-j||}{h}) Y_{jt}$$

decreases variance and increases Signal-to-Noise ratio (SNR)

reduce the number of independent decisions.

thresholds can be obtained by Random Field Theory (Adler 1987, 2000, Worsley 1994ff)

$$P(\max_{i}\bar{S}_{i} > \tau) \approx \sum_{d=0}^{3} R_{d}(V,h)\rho_{d}(\tau)$$

 $R_d(V,h)$  - d-dimensional resel count  $ho_d$ - d-dimensional Euler characteristic density.



decision using nonadaptive smoothing and thresholds given by Random Field Theory





#### Results of a voxelwise analysis,



Motor (finger tapping task)

# Interpretation of test results ?

- Poor signal to noise
- Activation or other sources of variation (motion artifact, physiological noise, other processes ?
- Color coded p-values
- Iow sensitivity <-> reduced spatial resolution
- Search for activated regions instead of activated voxel
- Reproducibility of results ??
  - large variability over repeated experiments (same subject)
  - Representativity for populations ?
  - Between subject variability
  - Group studies needed





fMRI experiments without external stimulus (resting state)

- looks for intrinsic brain activity
- first experiments by Biswal et al. (1995) observe patterns of spatial coherence between sensorimotor regions
- e.g. Zang & Raichle (2010) identify 7 mayor networks of regions that show spatially coherent activity
- larger studies identify up to 17 networks



Figure: Raichle, Brain Connectivity, 2011, Fig. 1D



# **ICA-Modell**



Independent component analysis (ICA)

- Observed signals  $x_1(t), x_2(t), \dots, x_p(t)$
- Assume these signals to be a linear combination of unknown souces  $s_1(t), s_2(t), \ldots, s_q(t)$

Model:

$$x_i(t) = \sum_{k=1}^{K} a_{ik} s_k(t) + \varepsilon_i(t) \quad i = 1, \dots, p$$

$$X = AS + E$$
(2)

- Goal: Estimate the mixing matrix  $A = (a_{ij})$  and the unknown source signals  $s_j(t)$
- Source separation or cocktail party problem



# **ICA in fMRI**



- Data:  $n_1 \times n_2 \times n_3 \times T$  values. Reorder as data matrix  $n \times T$
- Reduction of data matrix by Prewhitening and PCA, specification of number of sources K
- Search for spatial pattern in *S* (Spatial ICA)



Ylipaavalniemi and Vigário, Neuroimage 2008

- Decompose in temporal (A) and spatial S signals
- Solved by e.g. fastICA (Hyvärinen & Oja (2000))
- Some components *k* may model artifacts (interpretability !)





- C.F. Beckmann and S.M. Smith, Neuroimage 2005
- Generalization of ICA for group studies
- K subjects

Model:

$$X_{IK \times J} = (C | \otimes |A)S + E_{IK \times J} \tag{3}$$

$$(C|\otimes|A) = ((A\operatorname{diag}(c_1))^\top, \dots, (A\operatorname{diag}(c_K))^\top)^\top$$
(4)

- Structure of mixing matrix  $(C | \otimes | A)$  reflects the individual effects
- Common spatial structure in S





# 1200 Subjects

anatomical scans 0.7mm isotropic (T1/T2)

task based fMRI, 7 tasks 2mm isotropic
 (Working memomory, Gambling, Motor, Language, Social cognition, Relational processing, Emotion Processing)

- resting state fMRI 4 × 15min 2mm isotropic
  - diffusion weighted imaging 1.25mm isotropic,  $3 \times 90$  gradients

Information from these experiments is combined to obtain individual brain parcellations (node definitions) for all subjects

Literature:

- Special issue Neuroimage 2013
- Shen et al. Neuroimage 2013
- Finn et al., Nature neuroscience 2015





Brain parcellation, 268 functional regions, Shen 2013



- Finn 2015 defines general procedure for corresponding subject specific region definition
- regions should be used for node definition in group studies



- Selection of characteristic time series within regions
- leads to matrix  $Y = (y_{kt})_{k=1,K}^{t=1,T}$
- define network by empirical covariance matrix

$$\hat{\Sigma} = \left(\sum_{t} (y_{it} - y_{i.})(y_{jt} - y_{j.})\right)_{i,j=1,K}$$

- or regularized / thresholded estimate
- Task based fMRI depending on the goal:
  - Modeling and removal of expected hemodynamic response
  - Selection of characteristic (residual) time series within regions
- or
- Selection of nodes using functional regions associated with the tasks





# Electroencephalography (EEG)





Source: M. Lindquist, J. Hopkins Univ., Talk at SAMSI 2015

- high temporal resolution
- Iow spatial resolution
- Indirect measurement
- Source reconstruction problem
- Networks: coherence between spectra of recorded or reconstructed signals

Lit: Ombao and Van Bellegem (2008). Coherence Analysis: A Linear Filtering Point Of View. IEEE Transactions on Signal

Processing, 56(6), 2259-2266.





Assumptions:

$$Y_t \sim N_p(0, \Sigma), \qquad \Sigma = (\sigma_{ij})_{i,j=1}^p$$

Correlation between signals in nodes (regions) describes joint activity

$$R = (\rho_{ij})_{i,j=1}^p, \qquad \rho_{ij} = \frac{\sigma_{ij}}{(\sigma_{ii}\sigma_{jj})^{1/2}}$$





Assumptions:

$$Y_t \sim N_p(0, \Sigma), \qquad \Sigma = (\sigma_{ij})_{i,j=1}^p$$

Correlation between signals in nodes (regions) describes joint activity

$$R = (\rho_{ij})_{i,j=1}^{p}, \qquad \rho_{ij} = \frac{\sigma_{ij}}{(\sigma_{ii}\sigma_{jj})^{1/2}}$$

Partial correlations refer to joint activity not explained by intermediate effects

$$P = (\rho_{ij,k})_{i,j=1}^p, \qquad \rho_{ij,k} = \frac{\sigma_{ij} - \sigma_{ik}^T \Sigma_k^{-1} \sigma_{jk}}{((\sigma_{ii} - \sigma_{ik}^T \Sigma_k^{-1} \sigma_{ik})(\sigma_{jj} - \sigma_{jk}^T \Sigma_k^{-1} \sigma_{jk}))^{1/2}}$$
with  $k = (1 \dots n)/(ij)$ 

Precision matrices:  $\Omega = \Sigma^{-1} = (\omega_{ij})_{i,j=1}^p$ ,

Connection to partial correlations:  $ho_{ij,k} = -rac{\omega_{ij}}{(\omega_{ii}\omega_{jj})^{1/2}}$ 

- Pourahmadi, M.: Modeling covariance matrices: The GLM and regularization perspectives. Statist. Sci., 2011, 26, 369-87.



# Estimation of precision matrices (p << n)



Negative normal log-likelihood: 
$$Y_t \sim N_p(0, \Sigma), S = \frac{1}{T} \sum_{t=1}^T (Y_t - \bar{Y})(Y_t - \bar{Y})^T$$

$$\hat{\Omega} = \underset{\Omega}{\operatorname{argmax}} \log |\Omega| - tr(S\Omega), \qquad \hat{\Omega} = S^{-1}$$



- functional connectivity networks are hypothesized to be sparse
- p = 22, n = 178 → high variability of estimated correlations



## **Graphical LASSO**



## Regularization:

$$\hat{\Omega} = \operatorname*{argmax}_{\Omega} \log |\Omega| - tr(S\Omega) + \mathscr{P}_{\lambda}(\Omega)$$

Graphical LASSO:

$$\mathscr{P}_{\lambda}(\Omega) = \lambda \sum_{ij}^{p} |\omega_{ij}|$$

Literature:

- Meinshausen, N. & Bühlmann, P.: High-dimensional graphs and variable selection with the Lasso, Ann. Stat., 2006
- Friedman, J.; Hastie, T. & Tibshirani, R.: Sparse inverse covariance estimation with the graphical lasso, Biostatistics, 2008
- Levina, E.; Rothman, A. J. & Zhu, J.: Sparse estimation of large covariance matrices via a nested lasso penalty, Ann. Appl. Stat., 2008
- Rothman, A. J.; Levina, E. & Zhu, J.: Generalized Thresholding of Large Covariance Matrices, JASA, 2009
- Rothman, A. J., L. E. & Zhu, J.: Sparse multivariate regression with covariance estimation, JCGS, 2010
- Bien, J. & Tibshirani, R.: Sparse Estimation of a Covariance Matrix, Biometrika, 2011
- Rothman, A. J.: Positive definite estimators of large covariance matrices Biometrika, 2012
- Mazumder, R. & Hastie, T.: The Graphical Lasso: New Insights and Alternatives, Electr. J. Stat., 2012
- Mazumder, R. & Hastie, T.: Exact covariance thresholding into connected components for large-scale Graphical Lasso, JMLR, 2012



# **Graphical LASSO**



Regularization:

$$\hat{\Omega} = \operatorname*{argmax}_{\Omega} \log |\Omega| - tr(S\Omega) + \mathscr{P}_{\lambda}(\Omega)$$

Graphical LASSO:

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Solution for  $\lambda = .1$  function dpglasso from R-package dpglasso.

Problem: Produces a biased estimate !

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Regularization:

$$\hat{\Omega} = \underset{\Omega}{\operatorname{argmax}} \log |\Omega| - tr(S\Omega) + \sum_{ij}^{p} p_{\lambda}(\omega_{ij})$$

adaptive LASSO (Hui Zou):

$$p_{\lambda}(\omega_{ij}) = \lambda \frac{1}{\tilde{\omega}_{ij}^{\gamma}} |\omega_{ij}|$$

SCAD (Smoothly Clipped Absolute Deviation) ( Fan & Li (2001)):

$$p_{\lambda}(\omega_{ij}) = (\lambda I_{|\tilde{\omega}_{ij}| \le \lambda} + \frac{(a\lambda - |\tilde{\omega}_{ij}|)_{+}}{(a-1)} I_{|\tilde{\omega}_{ij}| > \lambda}) |\omega_{ij}|$$

Suggested parameters:  $\gamma = .5$ , a = 3.7.  $\tilde{\omega}$  are assumed to be consistent estimates.

Computations:

- Non-convex optimization problems
- can be approximated by iteration of graphical LASSO (with matrix penalty parameter)
- Zou, H.: The Adaptive Lasso And Its Oracle Properties, JASA, 2006, 101, 1418-1429
- Lam, C. & Fan, J.: Sparsistency and Rates of Convergence in Large Covariance Matrices Estimation, Ann. Stat, 2009
- Fan, J.; Feng, Y. & Wu, Y.: Network exploration via the adaptive LASSO and SCAD penalties, Ann. Appl. Stat, 2009
- Cai, T. T.; Liu, W. & Zhou, H.: Estimating sparse precision matrix: Optimal rates of convergence and adaptive estimation, Ann. Stat., 2014.
- Cai, T.; Liu, W. & Luo, X.: A Constrained I1 Minimization Approach to Sparse Precision Matrix Estimation, JASA, 2011





Regularization:

$$\hat{\Omega} = \operatorname*{argmax}_{\Omega} \log |\Omega| - tr(S\Omega) + \sum_{ij}^{p} p_{\lambda}(\omega_{ij})$$

$$p_{\lambda}(\omega_{ij}) = \lambda \frac{1}{\tilde{\omega}_{ij}^{\gamma}} |\omega_{ij}|$$

à



Parameters:  $\lambda = .1$ ,  $\gamma = .5$ .





Regularization:

$$\hat{\Omega} = \operatorname*{argmax}_{\Omega} \log |\Omega| - tr(S\Omega) + \sum_{ij}^{p} p_{\lambda}(\omega_{ij})$$

SCAD ( Fan & Li (2001)):

$$p_{\lambda}(\boldsymbol{\omega}_{ij}) = (\lambda I_{|\tilde{\boldsymbol{\omega}}_{ij}| \leq \lambda} + \frac{(a\lambda - |\tilde{\boldsymbol{\omega}}_{ij}|)_{+}}{(a-1)} I_{|\tilde{\boldsymbol{\omega}}_{ij}| > \lambda}) |\boldsymbol{\omega}_{ij}|$$



Parameters:  $\lambda = .1$ , a = 3.7.





Proposals (based on model selection criteria) with  $\Lambda = (\lambda_{ij})$ 

K-fold - Cross-validation

$$KCV(\Lambda) = \sum_{k=1}^{K} n_k (\log |\hat{\Omega}^{(-k)}(\Lambda)| - tr(S^{(k)}\hat{\Omega}^{(-k)}(\Lambda))$$

Generalized Cross validation (Dong & Wahba 1996, Lian 2011)

$$\begin{aligned} GACV(\Lambda) = &n(\log|\hat{\Omega}(\Lambda)| - tr(S\hat{\Omega}(\Lambda)) + \\ &+ \sum_{i=1}^{n} vec(\hat{\Omega}(\Lambda)^{-1} - y_i y_i^T)^T vec(\hat{\Omega}(\Lambda)(S^{(-i)} - S)\hat{\Omega}(\Lambda)) \end{aligned}$$

Bayes Information Criterion (BIC) (consistent !)

$$BIC(\Lambda) = -\log|\hat{\Omega}(\Lambda)| + tr(S\hat{\Omega}(\Lambda)) + k\frac{\log(n)}{n}$$

Suggestion: select maximum  $\lambda$  such that BIC slightly exceeds its minimal value.

- Lian, H.: Shrinkage tuning parameter selection in precision matrices estimation, J. Stat. Plan. Inf., 2011
- Chatterjee, A. & Lahiri, S. N.: Bootstrapping Lasso Estimators, JASA, 2011





SCAD (Fan & Li (2001)):

$$p_{\lambda}(\omega_{ij}) = (\lambda I_{|\tilde{\omega}_{ij}| \le \lambda} + \frac{(a\lambda - |\tilde{\omega}_{ij}|)_{+}}{(a-1)} I_{|\tilde{\omega}_{ij}| > \lambda}) |\omega_{ij}|$$

**Bayes Information Criterion (BIC)** 

$$BIC(\Lambda) = -\log|\hat{\Omega}(\Lambda)| + tr(S\hat{\Omega}(\Lambda)) + k\frac{log(n)}{n}$$



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Multiple precision Matrices:  $\Omega = (\Omega^{(1)}, \dots, \Omega^{(K)})$ 

$$\hat{\boldsymbol{\Omega}} = \operatorname*{argmax}_{\boldsymbol{\Omega}} \sum_{k=1}^{K} \log |\boldsymbol{\Omega}^{(k)}| - tr(\boldsymbol{S}^{(k)}\boldsymbol{\Omega}^{(k)}) + \mathscr{P}_{\boldsymbol{\lambda}}(\boldsymbol{\Omega})$$

Fused graphical LASSO:

$$\mathscr{P}_{\lambda}(\Omega) = \lambda_1 \sum_{k=1}^{K} \sum_{i \neq j} |\boldsymbol{\omega}_{ij}^{(k)}| + \lambda_2 \sum_{k' > k} \sum_{i,j} |\boldsymbol{\omega}_{ij}^{(k)} - \boldsymbol{\omega}_{ij}^{(k')}|$$

Group graphical LASSO:

$$\mathscr{P}_{\boldsymbol{\lambda}}(\Omega) = \lambda_1 \sum_{k=1}^{K} \sum_{i \neq j} |\boldsymbol{\omega}_{ij}^{(k)}| + \lambda_2 \sum_{i \neq j} \sqrt{\sum_{k=1}^{K} \boldsymbol{\omega}_{ij}^{(k)^2}}$$

Implementation: R-package(JGL)

- Tibshirani, R.; Saunders, M.; Rosset, S.; Zhu, J. & Knight, K.: Sparsity and smoothness via the fused lasso, JRSS B, 2005
- Yang, S.; Lu, Z.; Shen, X.; Wonka, P. & Ye, J.: Fused Multiple Graphical Lasso, see: http://people.math.sfu.ca/ zhaosong
- Danaher, P.; Wang, P. & Witten, D.: The joint graphical lasso for inverse covariance estimation across multiple classes, JRSS B, 2014





Multiple precision Matrices:  $\Omega = (\Omega^{(1)}, \dots, \Omega^{(K)})$ 

$$\hat{\boldsymbol{\Omega}} = \operatorname*{argmax}_{\boldsymbol{\Omega}} \sum_{k=1}^{K} \log |\boldsymbol{\Omega}^{(k)}| - tr(\boldsymbol{S}^{(k)} \boldsymbol{\Omega}^{(k)}) + \mathscr{P}_{\boldsymbol{\lambda}}(\boldsymbol{\Omega})$$

Fused graphical LASSO / SCAD:

$$\mathscr{P}_{\boldsymbol{\lambda}}(\Omega) = \sum_{k=1}^{K} \sum_{i \neq j} \lambda_{1ij} |\boldsymbol{\omega}_{ij}^{(k)}| + \sum_{k' > k} \sum_{i,j} \lambda_{2ij} |\boldsymbol{\omega}_{ij}^{(k)} - \boldsymbol{\omega}_{ij}^{(k')}|$$



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Source: Allen et al., Cerebral Cortex 2012.





## Learning paradigm:



Interest in changes of brain functionality due to learning



## Changes:

- Functional regions becoming active / inactive due to learning
- Changes in sets of regions that act coherently

Classical methods to detect these changes:

- Moving windows or comparison of first third and last third of time series
- Test if parameters / contrasts change over time
- Test if mean value of residuals changes over time
- Test if correlation / partial correlation matrices change over time







Test of stationarity without penalization:

$$H: \Sigma_t \equiv \Sigma \quad \forall t \in (h+1, n-h):$$

Use (log) Likelihood Ratio Test for

$$H_t: \Sigma_{t-} = \Sigma_{t+}$$

- Can be expressed in terms of eigenvalues  $l_1, \ldots, l_p$  of  $\hat{\Sigma}_t \hat{\Sigma}_{t+}^{-1}$
- $\Sigma_{t-}$  and  $\Sigma_{t+}$  estimated from left/right window of size h
- Test-Statistic:  $T(l_1, \dots, l_p) = -C_{h,p} \sum_{i=1}^p (\log(l_i) \log(1+l_i))$
- $\blacksquare \Rightarrow \mathsf{Curves} \ T(t,h), \quad t \in (h+1,n-h)$
- Distribution under Hypotheses H and  $H_t$  does not depend on  $\Sigma$  (as.  $\chi$ -square)
- Distribution under Hypothesis can be approximated by simulation  $\Rightarrow$  density  $d_h$

Problem: Test statistics undefined for h < p, highly variable if  $h \ge p$ Alternative proposal: Cai and Zhang, Inference for high-dimensional differential correlation matrices. JMVA 2016





Test of stationarity with penalization (GLASSO):

$$H: \Sigma_t \equiv \Sigma \quad \forall t \in (h+1, n-h):$$

Use (log) Likelihood Ratio Test for

$$H_t: \Sigma_{t-} = \Sigma_{t+}$$

- Can be expressed in terms of eigenvalues l<sub>1</sub>,..., l<sub>p</sub> of Σ̂<sub>t</sub>-Σ̂<sub>t+</sub><sup>-1</sup>
- **\Sigma\_{t-}** and  $\Sigma_{t+}$  estimated from left/right window of size *h*
- Test-Statistic:  $T(l_1, \dots, l_p) = -C_{h,p} \sum_{i=1}^p (\log(l_i) \log(1+l_i))$
- $\blacksquare \Rightarrow \mathsf{Curves} \ T(t,h), \quad t \in (h+1,n-h)$
- Distribution under Hypotheses H and  $H_t$  does depend on  $\Sigma$

Distribution of test statistic depends on unknown  $\Sigma$  and  $\lambda$ , may be approximated using permutation tests.

