Profiles of random trees

featuring infinitely-many-colour Pólya urns and "monkey" walks

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joint work with Jean-François Marckert (Bordeaux), Tim Rogers (Bath), and Gerónimo Uribe-Bravo (UNAM)

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Tree? Profile?



The profile of a tree τ is the probability measure

$$\Pi_{\tau} \propto \sum_{\nu \in \tau} \delta_{|\nu|},$$

where the sum's indices are the nodes ν of τ , and $|\nu|$ is the (graph) distance from ν to the root.

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Profiles of random trees

Describing the "shape" of random trees

One famous example: the Catalan tree (or Galton-Watson tree)

Pick a tree T_n uniformly among all *n*-leaf binary trees. What is its "typical shape" when $n \rightarrow \infty$?



Describing the "shape" of random trees

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- no scaling limit
- local limit: full infinite binary tree [Devroye '86]



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- local limit: full infinite binary tree [Devroye '86]
- profile: almost surely

$$\Pi_{\text{BST}_n}(\sqrt{2\log n} \cdot + 2\log n) \to \mathcal{N}(0, 1),$$

in the weak topology. [Chauvin, Drmota, Jabbour-Hattab '01]



Trailer of the talk

- The random recursive tree and its profile seen as an infinitely-many-colour Pólya urn. [with J.-F. Marckert]
- The weighted random recursive tree and its profile as a tool to study a "reinforced" random walk. [with G. Uribe-Bravo]
- Preferential attachment trees and their profile: an exciting open problem. [with T. Rogers]



























The profile is a Pólya urn: at each time-step, we

- pick a ball (node) uniformly at random, say, of colour (height) k
- add a new ball (node) of colour (height) k + 1.



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Don't panick!

Proving convergence in probability of the profile of RRT_n is actually the key to a general theory of infinitely-many-colour Pólya urns.

Convergence in probability of the profile of RRT_n

Theorem: [MM17]

When n goes to infinity, in probability

$$\Pi_{\text{RRT}_n}(\sqrt{\log n} \cdot + \log n) \to \mathcal{N}(0, 1),$$

for the weak topology on the space of probability distributions.

Ideas of the proof:

• Take U_n a node taken uniformly at random in RRT_n (cond. on RRT_n). It is easy to show that

$$\frac{|U_n| - \log n}{\sqrt{\log n}} \Rightarrow \mathcal{N}(0, 1), \text{ in distribution.}$$

This only implies that $\mathbb{E}\left[\Pi_{\mathbb{RRT}_n}(\sqrt{\log n} + \log n)\right] \rightarrow \mathcal{N}(0, 1).$

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Ideas of the proof:

• To prove the result, we need to take U_n and V_n two independent uniform nodes in RRT_n (cond. on RRT_n), and show that

$$\left(\frac{|U_n| - \log n}{\sqrt{\log n}}, \frac{|V_n| - \log n}{\sqrt{\log n}}\right) \Rightarrow (\Lambda_1, \Lambda_2),$$

where Λ_1 and Λ_2 are two independent standard Gaussians.

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The key arguments are

$$|U_n \wedge V_n| \to G \sim \text{Geom}(1/3)$$

- given G, the subtrees rooted at the two children of G
 - are i.i.d. random recursive trees and
 - have linear size in n.
Convergence of the profile of RRT_n

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Advertisement for [MM17]:

- Actually, using different techniques, we can prove convergence almost sure of the profile of the random recursive tree.
- One can define infinitely-many-colour urns as branching Markov chains on the random recursive tree.

Take: [Boyer et al. '14-'17]

- a "step" distribution on \mathbb{Z} , e.g. $\mathbb{P}(\xi = 1) = \mathbb{P}(\xi = -1) = 1/2$;
- a "run-length" distribution on \mathbb{N} , e.g. $\mathbb{P}(\omega = k) = (1/2)^k$.



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Main idea: drawing a position according to $\sum_{i=1}^{k} \mu_k$ is the same as

- drawing an integer I_k with probability $\mathbb{P}(I_k = i) \propto \omega_i$;
- 2 and then draw a position according to μ_{l_k} .



- draw a node (say node number *I_k*) of the WRRT at random with probability proportional to the weights (ω_i)_{i≥1};
- we add a child to this node, we draw a position X_k according to μ_{l_k}, and label the new node by the local time measure of a random walk starting at X_k and of length ω_k.

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 - the sequence of labels along each branch is a Markov chain (all of the same Kernel);
 - two distinct branches are independent after they branch.

We are interested in $\sum_{i=1}^{k} \mu_i / \sum_{i=1}^{k} \omega_i$ which is

- the local time of the monkey walk just before the *k* + 1-th relocation
- the distribution of the monkey walk at the *k* + 1-th relocation.

Understanding the Markov chain + the profile of the WRRT is enough to understand $\sum_{i=1}^{k} \mu_i$.

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Result and conjecture

Theorem [MUB++]

Let $(\omega_i)_{i\geq 1}$ be an i.i.d. sequence of weights of mean *m* and finite variance. Let $\mathbb{W}\mathbb{R}\mathbb{T}_k$ be the *k*-node weighted random recursive tree with weights $(\omega_i)_{i\geq 1}$. Then, $(\omega_i)_{i\geq 1}$ -a.s., in probability when $n \to \infty$,

$$\Pi_{\mathbb{W}\mathbb{R}\mathbb{T}_k}(\sqrt{\log n} \cdot + \log n) \to \mathcal{N}(0, 1).$$

Conjecture:

Let $(M_n)_{n\geq 0}$ be the monkey walk. In distribution when $n \to \infty$,

$$\frac{M_n - c_1 \log n}{\sqrt{c_2 \log n}} \Rightarrow \mathcal{N}(0, 1).$$

Further: can we prove convergence as a process to, say, a Brownian motion with random reinforced relocations?

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- At time 1, one node (the root).
- At time *n*, add the *n*th node in the tree: link it to a random node chosen with probability proportional to the degrees.



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Theorem [Katona '05]

Almost surely when $n \rightarrow \infty$ (for the weak topology),

$$\Pi_{\mathrm{BA}_n}\left(\sqrt{\frac{\log n}{2}} + \frac{\log n}{2}\right) \to \mathcal{N}(0,1).$$

- Fix $(X_n)_{n\geq 1}$ i.i.d. fitnesses.
- At time 1, one node (the root).
- At time *n*, add the *n*th node in the tree: link it to a random node chosen with probability proportional to the degrees **times the fitnesses**.

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Two competing dynamics: rich-gets-richer and fit-gets-richer.



Condensation

 μ = fitness distribution on [0, 1] X_i = fitness of node ν_i (i.i.d.)

Empirical fitness distribution:
$$\Xi_n = \frac{1}{n} \sum_{i=1}^n \deg(\nu_i) \delta_{X_i}$$
.

 Ξ_n converges (weakly) almost surely to



Conjecture:

In probability when $n \to \infty$ (for the weak topology): if (cond) fails, then $\Pi_{\text{BB}_n}(\sqrt{c \log n} + c \log n)$; if (cond) holds, then $\Pi_{\text{BB}_n}(\alpha_n + \beta_n) \to \mathcal{N}(0, 1)$ with $\alpha_n, \beta_n = o(\log n)$.

Conclusion

This talk in a nutshell:

- The profile of a random tree is a way to describe its shape particularly useful when there is no scaling/local limit, typically for "short fat" trees.
- Proving convergence in probability is done by taking two nodes at random in the tree and look at the joint distribution of their respective heights. (Almost sure convergence is much harder.)
- Convergence of the profile can be the first step towards studying more intricate objects such as the infinitely-many-colour urns and the monkey walk.
- Many fun open problems such as the BB-tree profile convergence.

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Thanks!!